

First Midterm Exam, Econ. 240B  
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**Instructions:** You have 110 minutes to complete this exam. This is a 30 point exam; all subsections of all questions have equal weight (5 points each). This is a closed book exam, but one sheet of notes is permitted. All needed statistical tables are appended. Please make your answers elegant – that is, clear, concise, and correct.

1. **True/False/Explain** (15 points): For **three** of the following four statements below, determine whether it is correct, and, if correct, explain why. If not, state precisely why it is incorrect and give a modification which is correct. **Answer only three questions;** if you answer more, only the **first three** answers will count in your score.

A. Suppose that an IV regression of  $y_i$  on a scalar endogenous regressor  $x_{i1}$  and a vector  $x_{i2}$  of exogenous regressors, using an instrument vector  $z_i$  that includes the  $x_{i2}$  components, yields a coefficient on  $x_{i1}$  of 2.2. If, instead,  $x_{i1}$  is taken to be the dependent variable, and an IV fit of  $x_{i1}$  on  $y_i$  and  $x_{i2}$  is calculated using the same instruments  $z_i$ , then the IV estimate of the coefficient on  $y_i$  will be positive.

B. In the linear model with a lagged dependent variable,  $y_t = x_t'\beta + \gamma y_{t-1} + \varepsilon_t$ , suppose the error terms are  $MA(1)$ , i.e.,  $\varepsilon_t = u_t + \theta u_{t-1}$ , where  $u_t$  is an i.i.d. sequence with zero mean, variance  $\sigma^2$ , and is independent of  $x_s$  for all  $t$  and  $s$ . For this model, the classical LS estimator will be inconsistent for  $\beta$  and  $\gamma$  when  $|\gamma| < 1$ , but an IV estimator using  $x_t$  and  $y_{t-2}$  as instrumental variables will consistently estimate these parameters.

C. For a balanced panel data regression model with individual fixed effects,  $y_{it} = x_{it}'\beta + \alpha_i + \varepsilon_{it}$  – where the  $\alpha_i$  are not assumed to be uncorrelated with  $x_{it}$ , but the error terms  $\varepsilon_{it}$  are i.i.d. and independent of  $\alpha_i$  and  $x_{it}$ , with  $E(\varepsilon_{it}) = 0$  and  $V(\varepsilon_{it}) = \sigma^2$  – suppose that only the number of time periods  $T$  tends to infinity, while the number of individuals  $N$  stays fixed. Then the “fixed effect” estimator for  $\beta$  will be consistent as  $T \rightarrow \infty$  provided the regressors and individual indicator variables are not asymptotically multicollinear. Furthermore, if  $\hat{\sigma}^2 = (NT)^{-1} \sum_i \sum_t (y_{it} - \hat{\alpha}_i - x_{it}'\hat{\beta}_{LS})^2$  is the (biased) LS estimator of  $\sigma^2$ , then the usual LS formulae for the standard errors of  $\hat{\beta}_{LS}$  (replacing the unknown  $\sigma^2$  by  $\hat{\sigma}^2$ ) will be asymptotically valid.

D. By the so-called "Delta Method", if  $\hat{\theta}$  is root- $n$  consistent and asymptotically normal for a vector parameter  $\theta_0$ , then the difference between the squared length of  $\hat{\theta}$  and the squared length of  $\theta_0$ , when multiplied by the square root of the sample size, will generally have a limiting normal distribution.

2. (5 points) Suppose a dependent variable  $y_i$  and two (scalar) regressors  $x_i$  and  $z_i$  satisfy a random coefficients model

$$y_i = \alpha_i + \beta_i x_i + \gamma_i z_i, \quad i = 1, \dots, N,$$

where the coefficients  $(\alpha_i, \beta_i, \gamma_i)$  are assumed to be i.i.d. and independent of  $x_i$  and  $z_i$ . In this framework, under the null hypothesis  $H_0 : Var(\beta_i) = 0 = Var(\gamma_i)$ , the mean values  $\beta = E(\beta_i)$  and  $\gamma = E(\gamma_i)$  can be estimated by a least-squares regression of  $y_i$ ; in turn, this null hypothesis can be tested using the  $R^2$  from a least-squares regression of the squared LS residuals  $\hat{e}_i^2 = (y_i - \hat{\alpha} - \hat{\beta}x_i - \hat{\gamma}z_i)^2$  on functions of the regressors.

Given a sample of size  $N = 500$ , derive the algebraic form of all of the regressors in this "squared residual regression", and give a numerical value for the critical value  $C$  for an (asymptotic) 5% test of homoskedasticity using the second-stage  $R^2$ . i.e., the value for which  $H_0$  will be rejected if  $R^2 > C$  with asymptotic size 5%.

3. (5 points) A feasible GLS fit of the generalized regression model with  $K = 3$  regressors yields the estimates  $\hat{\beta} = (2, -2, -1)$ . where the GLS covariance matrix  $V = \sigma^2[X'\Omega^{-1}X]^{-1}$  is estimated as

$$\hat{V} = \begin{bmatrix} 2 & 0 & 1 \\ 0 & 1 & 0 \\ 1 & 0 & 1 \end{bmatrix},$$

using consistent estimators of  $\sigma^2$  and  $\Omega$ . The sample size  $N = 403$  is large enough so that it is reasonable to assume a normal approximation holds for the GLS estimator.

Use these results to test the null hypothesis  $H_0 : \beta_1^2 + \beta_2^2 + \beta_3^2 = 1$  at an asymptotic 5% level.

4. (5 points) If  $y_t$  is an  $MA(1)$  process with zero mean, i.e., if

$$y_t = \varepsilon_t + \theta\varepsilon_{t-1}, \quad \varepsilon_t \sim WN(\sigma^2),$$

and if  $\gamma(s) = Cov(y_t, y_{t-s})$  is the autocovariance function and  $\rho(s) = \gamma(s)/\gamma(0)$  is the autocorrelation function of  $\{y_t\}$ , show that

$$-1 < c^L \leq \rho(1) \leq c^U < 1,$$

i.e., the first autocorrelation is strictly bounded away from  $-1$  and  $1$ , by calculating the maximum and minimum values  $c^U$  and  $c^L$  of  $\rho(1)$  over all possible  $\theta$ .